# Finding the optimal sparse, overcomplete model for natural images by model selection

# Pietro Berkes, Richard Turner and Maneesh Sahani {berkes,turner,maneesh}@gatsby.ucl.ac.uk

Gatsby Computational Neuroscience Unit, Alexandra House, 17 Queen Square, London.

# Abstract

Computational models of visual cortex, and in particular those based on sparse coding, have enjoyed much recent attention. Despite this currency, the question of how sparse or how over-complete a sparse representation should be, has gone without principled answer. Here, we use Bayesian model-selection methods to address these questions for a sparse-coding models based on a Student-t prior and on a Gaussian scale mixture model with uniform prior on precision. We find that natural images are indeed best modeled by extremely sparse distributions, although for these priors, the associated optimal basis size is only modestly over-complete (Berkes et al., 2008).

# Linear, sparse coding model

$$\mathbf{y}_t = \sum_i \mathbf{g}_i x_{i,t} + \epsilon_t, \quad p(x_{i,t} | \alpha) = p_{\text{sparse}}(\alpha)$$
(Olshausen & Field, 1996, 1997)

Overcomplete case: # latent variables > # input dimensions An overcomplete 1D model with 2 components:



The posterior distribution in very sparse, overcomplete models are complex and multimodal.

# Addressed questions

## How sparse? Which family of distributions?

• How overcomplete? (overfitting) (see also Olshausen and Millman, 2000)

• One might expect a tradeoff between sparseness and overcompleteness

### Model selection

• One possibility is to implement different models and find the one which is most "similar" to visual processing

<ul> <li>Bayesian pers</li> </ul>	pective:				
<ul> <li>compare</li> </ul>	marginal	likelihood	of	the	model
$\underline{p}(\mathcal{M}_1, \Xi_1   Y)$	$\underline{p}(Y \mathcal{M}_1,$	$\Xi_1$ ) $P(\mathcal{M}_1,$	$, \Xi_1)$	P(Y .	$\mathcal{M}_1, \Xi_1)$
$\overline{p(\mathcal{M}_2, \Xi_2 Y)}$	$- p(Y \mathcal{M}_2,$	$\Xi_2$ ) $P(\mathcal{M}_2)$	$, \Xi_2)$	$\overline{P(Y)}$ .	$\mathcal{M}_2, \Xi_2)$
<ul> <li>automatic Occam's razor</li> </ul>					
<ul> <li>natural if</li> </ul>	hypothesis	s is that	the	visual	system
implements	an optimal g	generative r	nodel		
References:					

P. Berkes, R. Turner, and M. Sahani, On sparsity and overcompleteness in image models. In Advances in Neural Information Processing Systems, 20, 2008. M.J. Beal. Variational Algorithms for Approximate Bayesian Inference. PhD thesis, Gatsby Computational Neuroscience Unit, University College London, 2003. C.M. Bishop. Variational principal components. In ICANN 1999 Proceedings, 509–514, 1999. B.A. Olshausen and D.J. Field. Emergence of simple-cell receptive field properties by learning a sparse code for natural images. *Nature*, 381(6583):607–609, 1996. B.A. Olshausen and D.J. Field. Sparse coding with an overcomplete basis set: A strategy employed by V1? Vision Research, 37:3311-3325, 1997. B.A. Olshausen and K.J. Millman, Learning sparse codes with a Mixture-of-Gaussians prior. In Advances in Neural Information Processing Systems, 12, 2000. R.M. Neal. Annealed importance sampling. Statistics and Computing, 11:125–139, 2001.

# Model selection in practice

•Strategy: • Use Automatic Relevance Determination (ARD) prior, Bayes Expectation Maximization (VBEM) to Variational determine the posterior over parameters and the overcompleteness • Unfortunately VBEM is biased (cannot use the free energy) • Compute the likelihood of the learned model using Annealed Importance Sampling (AIS)

• Gaussian prior over the components that favors small weights; hyperprior over the precisions to keep the prior uninformative

• Start with excess of components, let the inference process prune the weights which are unnecessary Learning using VBEM

# Why VBEM is biased

 $\log p(Y|\mathcal{M}, \mathbb{R})$ 

The free energy bound is tightest where q(V,Theta) is a good match to the true posterior. At high sparsities, the true posterior is multimodal and highly non-Gaussian. At low sparsities, the true posterior is Gaussian-like and unimodal. q(V,Theta) is always unimodal. There is an additional bias toward complete solutions, due to the fact that the posterior of an overcomplete solution is more multimodal and thus less Gaussian. This is investigated in a set of simulations with artificial data. The simulations confirm that the bias exist, but the solution found are still considerably more overcomplete than those found in the main simulations.

# Annealed Importance Sampling (Neal, 2001)

 $p_j(x) = p_0(x)$  $0 = \beta_N >$ 

Idea 2: Importance sampling: View annealing as defining an importance sampling distribution over  $(x_0, \ldots, x_{N-1})$ 

Guarantees asymptotic correctness.



# **ARD** (Bishop, 1999; Beal, 2003)

$$p(\mathbf{g}_k|\gamma_k) = \mathcal{N}_{\mathbf{g}_k}(\mathbf{0}, \gamma_k^{-1})$$
$$p(\gamma_k) = \mathcal{G}_{\gamma_k}(\theta_k, l_k) .$$

 $\Xi$ : hyperparameters Z: latent variables  $\Theta$  : parameters

$$\Xi) \ge \int dV d\Theta \ q(V, \Theta) \log \frac{p(Y, V, \Theta | \mathcal{M}, \Xi)}{q(V, \Theta)} =: \mathcal{F}(q(V, \Theta))$$
$$= \log p(Y | \mathcal{M}, \Xi) - KL(q(V, \Theta) || p(V, \Theta | Y))$$

## Idea 1: Simulated annealing

$(x)^{\beta_j} p_N(x)^{1-\beta_j}$	$p_N(x)$ : prior distribution
$\ldots > \beta_0 = 1$	$p_0(x)$ : unnormalized posterior distributio

 $p_N \sim x_{N-1} \xrightarrow{T_{N-1}} x_{N-2} \xrightarrow{T_{N-2}} \dots \xrightarrow{T_1} x_1 \xrightarrow{T_0} x_0$  $x_{N-1} \underset{\tilde{T}_{N-1}}{\leftarrow} x_{N-2} \underset{\tilde{T}_{N-2}}{\leftarrow} \dots \underset{\tilde{T}_1}{\leftarrow} x_1 \underset{\tilde{T}_0}{\leftarrow} x_0 \sim p_0$ 

# Artificial data



# Image-like data

### Natural data



